

Age-specific incidence of inherited versus sporadic cancers: A test of the multistage theory of carcinogenesis

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Knudson [Knudson, A. G. (1971) *Proc. Natl. Acad. Sci. USA* 68, 820–823] suggested that progression of retinoblastoma follows from two mutational events. Individuals who inherit one mutated gene copy should follow an age-onset pattern set by only a single rate-limiting step for transformation, whereas normal individuals should follow an age-onset pattern set by two rate-limiting events. Knudson's analysis of inherited and sporadic cases of retinoblastoma supported this prediction. However, retinoblastoma has a peculiar age-onset pattern concentrated in early life, because the retinal tissue completes most of its cell division by 5 years of age. Here, I compare age-specific incidences of inherited and sporadic forms of colon cancer, a much more typical form of human cancer. My simple mathematical analysis based on multistage theory explains the observed differences in age-onset patterns between inherited and sporadic cases. I also analyze recent retinoblastoma data and provide a mathematical analysis and interpretation. My analysis supports Knudson's two-hit theory but is much simpler and easier to understand than the original mathematical theory, which was based on a complicated model of cell division in the retina. My simpler theory for retinoblastoma makes clear the common basis for understanding multistage progression in tissues as different as the retina and colon.

cancer epidemiology | inherited predisposition | retinoblastoma

The age of cancer onset follows remarkably simple patterns. The most common cancers arise in epithelial tissue late in life, increasing in incidence with roughly the fifth or sixth power of age (1). By contrast, the incidences of many childhood cancers decline steadily with age as the particular tissues in which they occur slow their rate of cell division with advancing development (2).

The simple regularity in age-specific incidence seems at odds with the great complexity and apparent randomness of environmental and genetic aspects of carcinogenesis. Yet, where one observes simple patterns, one suspects a simple process at some level of explanation.

For the past 50 years, Armitage and Doll's (1) multistage theory of cancer has provided the conceptual foundation for simple explanations of the observed age-specific incidence patterns. The multistage theory suggests that normal tissues are transformed into cancerous ones by means of a series of discrete stages. The stages may be somatic mutations, broad genomic rearrangements, or changes in tissue interactions and environment.

The precise nature of the stages does not affect the predictions of the multistage theory. What matters is that individuals, as they age, move stochastically through the various stages of transformation. Then, at any particular age, there is a regular probability distribution of individuals who have progressed to particular precancerous stages or all of the way to the final, malignant stage. What happens to an individual is highly random and cannot be predicted. By contrast, a distinct and predictable pattern emerges at the population level.

Here, I test the predictions of multistage theory by comparing the age-specific incidences of cancer in individuals who inherit a

predisposing genetic mutation versus those individuals who do not inherit such a mutation. It was Knudson's (3) great insight to compare incidences in inherited and noninherited forms of retinoblastoma to test the predictions of multistage theory.

Methods

I extended Knudson's general approach in two ways. First, I compared age-specific incidences in inherited and noninherited forms of colon cancer. I also developed the specific mathematical prediction for this comparison that follows from multistage theory. This analysis and comparison to the data add to the subject because the only similar comparison, which was on retinoblastoma, was for a cancer with unusually simple genetics and a peculiar pattern of childhood onset. By contrast, colon cancer has the age-incidence profile of the common epithelial cancers that account for the great majority of all human cancers.

Second, I analyzed recent retinoblastoma data, again comparing the age-specific incidences for inherited and noninherited forms. I developed a very simple multistage theory that explains the observed differences. Knudson's original papers chose the right comparisons to illuminate the theory and correctly interpreted the data. But the first mathematical analysis (3) did not account for the patterns of cell division in retinal development, and the later papers (4, 5) accounted for cell division but were so mathematically complex that the elegant insight of comparing inherited and noninherited cases was lost in mathematical detail and technical procedures for fitting the model to the data. By contrast, I show how a simple quantitative analysis can bring out Knudson's insight in a more accessible and profound way.

Results and Discussion

Colon Cancer. Individuals who inherit one mutated copy of the *APC* gene almost invariably develop multiple colon tumors by midlife, causing a disease known as familial adenomatous polyposis (FAP) (6). In terms of multistage models, individuals with an inherited *APC* mutation begin life one stage further along than do normal individuals.

Fig. 1A shows the age-specific incidence for individuals with inherited FAP or noninherited (sporadic) colon cancer. On log-log scales, both inherited and sporadic forms show approximately linear increases in incidence with age. The two lines are nearly parallel, with the inherited form occurring at an incidence rate 3–4 orders of magnitude greater than that for sporadic cases. Note that the incidence rate for inherited cases is given relative to the population of individuals carrying the inherited mutation, whereas the sporadic incidence is given relative to the

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Abbreviations: FAP, familial adenomatous polyposis; SEER, Surveillance, Epidemiology, and End Results.

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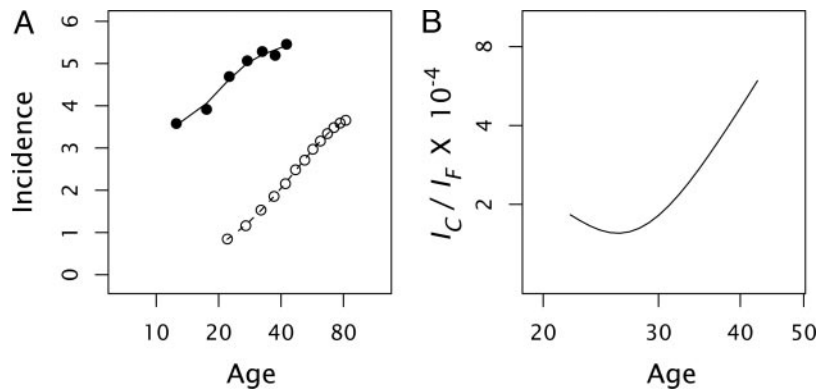


Fig. 1. Age-specific incidence of inherited and sporadic colon cancer. (A) Inherited colon cancer (FAP) caused by mutation of the *APC* gene (●) and sporadic cases (○) per 10^6 population, shown on a \log_{10} scale. I calculated FAP incidence by analyzing the age distribution of cases combined for males and females as summarized by Ashley (7) from data originally presented by Veale (8). Mutated *APC* alleles have very high penetrance for FAP, so the incidence at each age can be measured as the number of cases per year divided by the number of individuals who had not developed the disease in earlier years but who do eventually develop the disease. For the sporadic form, I used the incidence of colorectal cancers from the Surveillance, Epidemiology, and End Results (SEER) database combined for white males and females from the period 1973–1977 (<http://seer.cancer.gov>). (B) Ratio of sporadic colon cancer incidence (I_C) to inherited FAP incidence (I_F) at each age by using the data in A. This curve is obtained by taking the ratio of the fitted curves in A.

population of noncarriers. Thus, the frequency of carriers does not influence the rate curves.

Fig. 1B plots the ratio of sporadic incidence to inherited age-specific incidence, I_C/I_F . This ratio increases ≈ 3 -fold with age, varying between ≈ 2 and 6×10^{-4} .

What does multistage theory predict for the ratio of sporadic to inherited incidence? It is possible to construct elaborate multistage models based on assumptions about tissue architecture, cell division rate, clonal expansion of mutated cells, and so on. Such mathematical theories can provide significant insight. However, it is difficult to test such models directly, because there are always other factors not included in the model, so any fit between theory and data is as likely to be a matter of fortuitous fitting of parameters as it is to be a matter of capturing the essential processes that explain most of the observed variation.

For these reasons, I prefer to use the simplest theory. Simple theory makes clear predictions about how observations should change when underlying parameters change. For example, how much greater should the incidence of inherited cancers be relative to sporadic cases? How should the ratio of sporadic to inherited cases change with age? If we can consistently succeed with such simple predictions, then perhaps we have captured the major features that control incidence rates. Of course, it is always possible that we could be right for the wrong reasons, that is, other models may also explain the data. But at least we should start with a simple approach and see how well we can succeed with simple explanations.

The simplest multistage model assumes that there are n stages of cancer development (1). The tissues of a normal individual begin life in stage 0 and slowly make stochastic transitions through the stages as age increases. In particular, suppose the rate of transition between stages is u . For individuals in stage $n - 1$, the rate of incidence will be u , the rate of the final transition. This rate is independent of time, because the transition hits randomly and so is equally likely to occur in any particular year. Thus, the incidence rate at age t is the probability of being in stage $n - 1$, multiplied by the rate u . If both inherited and noninherited cases must make the same final transition, then the relative incidence rate, R , of noninherited and inherited cases is the relative probability of being in stage $n - 1$ at age t .

The probability of any single transition over a period of t years is ut , the rate of transition multiplied by the total time elapsed. By using a widely known result from probability theory based on the gamma distribution, the probability at time t of k events

happening in a particular order, each event having probability ut , is approximately $P_k = (ut)^k/k!$, as long as ut is not too close to 1.

In individuals born without a mutation, a cell or tissue moving from stage 0 to stage $n - 1$ requires $n - 1$ transitions. By contrast, for individuals who inherit a mutation, only $n - 2$ transitions are required. Thus, the relative probability of being in stage $n - 1$ for the noninherited versus inherited group is the relative incidence, R , which is

$$R = \frac{P_{n-1}}{P_{n-2}} = \frac{ut}{n-1}.$$

If transitions occur as somatic mutations, then the transition rate per year is the mutation rate per cell division, v , multiplied by the number of cell divisions per year, C . Few attempts have been made to measure the somatic mutation rate per gene per cell division. Yeast provide a convenient model of single eukaryotic cells. For yeast, the mutation rate has been estimated at $10^{-7} - 10^{-5}$ (9, 10). In mice, Kohler *et al.* (11) estimated the frequency of somatic mutations at 1.7×10^{-5} . There are $\approx 10^1$ to 10^2 cell divisions back to the embryo, so this study suggests a somatic mutation rate per cell division on the order of 10^{-7} to 10^{-6} . I use the approximate value of 10^{-6} per gene per cell generation, but this is a very rough estimate at present.

Colon epithelial tissue renews itself continuously throughout life. The surface tissue turns over every few days, and stem cells that ultimately renew the tissue probably divide at least once per week or ≈ 50 times per year. For the number of stages, I use the simplest estimates from epidemiological data, which suggest that the number of stages in colon cancer progression is around $n = 6$ (1); other specialized models have, for example, put the number at $n = 4$ (12). All of these numbers are provisional, but they allow us to predict that the ratio of sporadic to inherited incidence rates should be roughly

$$R = \frac{ut}{n-1} \approx 10^{-5}t.$$

The data for inherited FAP and sporadic cases can be compared on the range $t = 20$ – 40 , so R is predicted to increase over the range 2 – 4×10^{-4} . Fig. 1B shows that the ratio of incidences is of the predicted magnitude and increases with age, although the increase with age is slightly greater than predicted. The nonlinearity between ages 20 and 30 may arise from statistical fluctu-

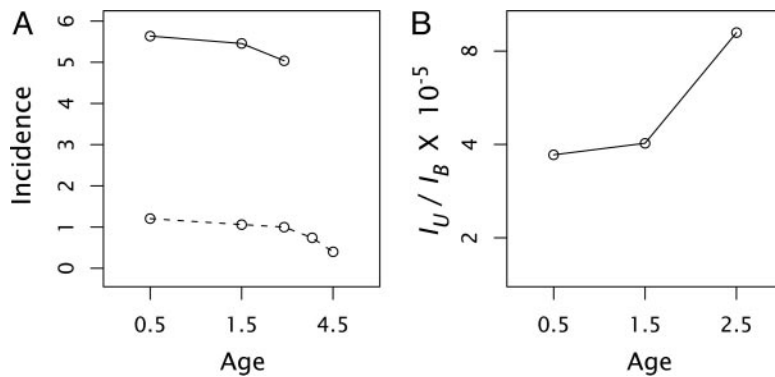


Fig. 3. Age-specific incidence of retinoblastoma. (A) Bilateral (solid line) and unilateral (dashed line) cases of retinoblastoma per 10^6 population, shown on a \log_{10} scale. (B) Ratio of unilateral (I_U) to bilateral (I_B) incidence at each age by using the data in A. For the inherited form, I used 221 reported bilateral cases taken directly from the SEER database (<http://seer.cancer.gov>) for 1973–2001. To estimate age-specific incidence, I assumed that 65% of carriers eventually developed bilateral tumors based on the estimated penetrance for bilateral retinoblastoma given by Knudson (3). The incidence in each year is approximately the fraction of cases in that year divided by the fraction of individuals who had not developed the disease in earlier years. For the sporadic form, I used the reported incidence of unilateral cases in Young *et al.* (16), which is also from the SEER database. However, the SEER data do not differentiate between sporadic and hereditary unilateral cases. Based on data from Knudson (3), $\approx 75\%$ of unilateral cases are sporadic cancers and $\approx 25\%$ arise from carriers who inherit a mutation.

I advocate theory more closely matched to Knudson's original insight and what one can realistically infer given the nature of the data. According to Knudson's theory, bilateral tumors arise from single hits to somatic cells with an inherited mutation. The rate at which a hit occurs in the developing retina at a particular age depends on many factors, including the number of target cells and the rate of cell division. But good estimates for those factors are not available, so, instead, the observations for bilateral cases at different ages can be used to estimate the rate at which a somatic mutation occurs in the tissue at a particular age, subsuming all of the details that together determine that rate. In particular, the estimates for age-specific bilateral incidence are taken as the estimates for the rate at which second hits occur in the tissue at a particular age. Clearly, this estimation procedure simplifies the real process; for example, bilateral cases require at least one hit in each eye. However, the rate of two second hits leading to bilateral cases is fairly high at $\approx 10\text{--}30\%$ (Fig. 3A); thus, the rate of one second hit is about the same order of magnitude as the rate of two second hits. So I proceeded with the simple approach that $I_B(t)$, the incidence of bilateral cases at age t , provides a rough estimate of the rate of second hits to the tissue at age t .

The incidence of unilateral cases can be written as

$$I_U(t) = f(t)I_B(t),$$

where $f(t)$ is the fraction of somatic cells at age t that carry one somatic mutation, and $I_B(t)$ is approximately the rate at which the second hit occurs. Knudson's insight was to compare the incidences of unilateral and bilateral cases, so we could study the ratio of unilateral to bilateral incidence at each age

$$\frac{I_U(t)}{I_B(t)} = f(t).$$

In words, the ratio of unilateral to bilateral rates should be roughly $f(t)$, the fraction of cells at time t that carry the first hit in individuals that do not inherit a mutation. For example, if $f(t) = 1$, then all somatic cells have a first mutation, and the susceptibility is the same as for inherited cases. If $f(t) = 0.5$, then one-half of the somatic cells carry a first hit, and the susceptibility is one-half that of individuals who inherit the mutation.

The expected number of somatic mutational events suffered by an allele in a particular cell is the mutation rate per cell

division, ν , multiplied by the number of cell divisions going back to the embryo. Let the number of cell divisions at age t be $C(t)$, so that $\nu C(t)$ is the expected number of mutational events. For most assumptions, $\nu C(t) < 1$, so we can take $\nu C(t) = f(t)$ as the fraction of cells at time t that carry a somatic mutation, and thus

$$\frac{I_U(t)}{I_B(t)} = \nu C(t).$$

As discussed earlier, I used the approximate somatic mutation rate per cell division of $\nu \approx 10^{-6}$. The number of cell divisions, $C(t)$, is roughly in the range of 15–40, because there are probably $\approx 15\text{--}25$ cell divisions before the start of retinal development, and it takes ≈ 15 cellular generations in the retina to make the $e^{15} \approx 10^6$ to 10^7 cells in the fully developed retina. Thus, $I_U(t)/I_B(t) \approx 10^{-4}$ to 10^{-5} , and this ratio may increase by a factor of about two during early childhood as $C(t)$ increases from $\approx 15\text{--}25$ at the start of retinal development to $\approx 30\text{--}40$ in the final cellular generations in the retina.

These rough calculations lead to two qualitative predictions. First, the ratio of unilateral to bilateral age-specific incidence should be $\approx 10^{-4}$ to 10^{-5} . Second, the ratio of unilateral to bilateral incidence should approximately double with age over the period of retinal growth as the number of cellular generations, $C(t)$, increases with time.

Fig. 3B shows that the ratio of unilateral and bilateral incidence is in the predicted range of 10^{-4} to 10^{-5} , roughly the somatic mutation rate multiplied by the number of cellular generations. This ratio approximately doubles from the earliest age of 0–1 to the latest age of 2–3 at which sufficient numbers of bilateral cases occur to estimate incidence rates. The increase of this ratio supports the prediction that unilateral incidence increases relative to bilateral incidence as the number of cellular generations increases.

Conclusions

Knudson's insight was to test multistage theory by comparing the age-specific incidences of inherited and sporadic forms. The original mathematical theory for retinoblastoma was precise (5) but so complicated that it did not provide a clear and accessible test of the simple predicted comparison. In this paper, I developed simple mathematical analyses to emphasize the comparison between inherited and sporadic age-specific incidences in colon

cancer and retinoblastoma. My mathematical models explain the observations and show the common aspects between these very different cancers.

For complex biological problems, such as cancer, comparative predictions and simple mathematical analyses provide more insight than complex and mechanistically more detailed mathematical models. The complex models require fitting several parameters from the data, and such parameter estimates based

on highly variable data rarely provide accurate measures of the true processes. By contrast, comparative predictions, such as Knudson's analysis of sporadic versus inherited retinoblastoma, show how a simple difference in biological process can explain major differences in observed outcomes.

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